

Household Food Insecurity Is Not Associated with BMI for Age or Weight for Height among Brazilian Children Aged 0–60 Months

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Abstract

We examined the association between Household Food Insecurity (HFI), weight for height z-score (WHZ) and Body Mass Index for age z-score (BMI-Z) in a representative sample of children 0–60 months of age ($n=3,433$) in five Brazilian geographical regions. Data were derived from the 2006–07 Brazilian Demographic and Health Survey. HFI was measured with the Brazilian Food Insecurity Scale. Associations were estimated using multiple linear regression models (β coefficients and 95% CI) taking into account the complex sampling design. Interaction terms between HFI and geographical region and HFI and child sex and child age were assessed. The weighted prevalence of any level of HFI was 48.6%. Severe food insecurity was more prevalent among children from the North region (16.8%), born from mothers with <4 years of schooling (15.9%) and those from families with ≥ 3 children (18.8%). The interaction between HFI and geographical region was non-significant for BMI-Z ($P=0.119$) and WHZ ($P=0.198$). Unadjusted results indicated that HFI was negatively associated with BMI-Z (moderate to severe HFI: $\beta = -0.19$, 95% CI: $-0.35 - -0.03$, $P=0.047$), and WHZ (moderate to severe HFI: $\beta = -0.26$, 95% CI: $-0.42 - -0.09$, $P=0.009$). Estimates lost significance after adjustments for key confounders such as mothers' skin color, mothers' years of schooling, place of household, household income quartiles, mothers' smoking habit, mothers' marital status, number of children 0–60 months in the household, and birth order. HFI is unrelated to weight outcomes among Brazilian children 0–60 months.

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Introduction

Women living in households with food insecurity (HFI) seem to be at increased risk of obesity [1,2]. However, among children, support for an association between HFI and weight status has been mixed [3–5]. A small number of studies have found that children living in food insecure households are more likely to be overweight or obese [6–9], however other studies have not found an association between HFI and weight status [10,11] or have found that HFI is associated with lower risk of pediatric obesity [12,13]. Some studies that reported that HFI is associated with a higher risk of overweight found that this association was present only among girls but not among boys [6,7]. Most of these studies were carried out in high-income countries and combined children with adolescents in the analysis [2,4,9]. In low and middle-income countries results are also mixed [14–19].

Our research group has previously documented associations between HFI and obesity risk among adult and adolescent women using data from the latest Brazilian Demographic and Health Survey (DHS) [1,20]. However the relationship between HFI and the risk of childhood obesity among different sex and age-subgroups of Brazilian children remains to be elucidated. The answer to this question is essential to understand if there is a relationship between HFI and obesity risk across different stages of the life cycle.

The objective of this paper is to examine the association between HFI, weight for height z-score (WHZ) and Body Mass Index for age z-score (BMI-Z) in a representative sample of children 0–60 months of age in five Brazilian geographical regions after adjusting for key potential confounders. Because Brazil is a highly diverse country we also examined the interaction between

HFI and geographical region on each of these anthropometric outcomes.

Methods

Ethics Statement

The project was approved by the Research Ethics Committee of the Sexually Transmitted Diseases/AIDS Reference and Training Centre of the Health Secretariat of the state of São Paulo.

Study design and sampling

The data were derived from the third wave of the DHS, conducted in Brazil, in 2006–07. DHS was a population-based cross-sectional study with a nationally representative sample of women of reproductive age and mothers of children aged 0–60 months. DHS included both household- and individual- but no community- level variables, and has a complex sampling design. The sample units were selected in two stages. The first stage involved selecting the census sectors, which were the primary sampling units. The second stage involved selecting the survey participants within each household. Ten sampling strata were defined based on a combination of urban *vs.* rural areas and the five Brazilian geographical regions. The respondents' sampling weights were derived from the household sampling weights and took into account the probability that there may be more than one eligible woman per household. The weights were adjusted for non-response within households and were calibrated based on official population estimates released by Brazilian Institute of Geography and Statistics.

Structured questionnaires were applied to the mothers or caregivers of the children through in-person interviews and children's anthropometric measures were taken. Data collected included socioeconomic status, lifestyle variables, reproductive history, and HFI.

Anthropometric data and study sample

Weight and length or heights of all eligible children in the selected households were measured according to the recommendations of the World Health Organization [21]. These measurements were conducted twice for each subject, and the mean value was calculated. Weight was measured using an electronic scale (Dayhome[®]) with a 100 g precision, which was calibrated at the beginning and at the end of each working day. Length or height was measured using stadiometers made locally for the purpose of the study with 1 mm precision that were calibrated at the beginning and at the end of each working day. BMI was calculated as weight (kg)/length or height (m²). The anthropometric assessments were performed by evaluators who had completed high school, and who received a 32-hour training course on child anthropometry, focusing mainly on standardizing length and height measurements. The evaluators' performance was assessed against a gold-standard anthropometrist, and only those demonstrating adequate accuracy and precision participated on the data collection.

The initial analytical sample was formed by 5,146 children, of both sexes aged 0–60 months, living in 4,108 households. The analytical sample considered for these analyses included the 3,527 subjects that had available data for all key variables (age, sex, weight, length or height and HFI), and covariates. The variables with the most missing data were weight and height/length ($n = 1,184$) and family income. We excluded those individuals classified as outliers ($n = 94$), assuming that their anthropometric measures were physiologically implausible, using the following criteria: Height for age z-score (HAZ) < -5 and > +3, Weight for

age z-score (WAZ) < -5 and > +5 and WHZ < -4 and > +5 (21). Although HAZ and WAZ were not directly measured, it was necessary to identify outliers for these indicators based on physiologically implausible limits for weight or height measures, as both are needed for estimating the study outcomes (BMI-Z and WHZ).

The final sample comprised 3,433 children aged 0–60 months. The age categorization (0–24; 25–60 months) was implemented in order to group children with similar anthropometric data collection protocols for length (<24 months) or height (≥ 24 months), and also due to expected differences in growth patterns among these age groups.

Outcome and independent variables

In our analyses we included two different anthropometric outcomes, first to check results consistency between outcomes and second to fully examine the association between HFI and different measures of weight status. BMI-Z and WHZ were calculated according to the World Health Organization (WHO) growth reference standards [22]. BMI-Z can measure either excess or insufficient weight for a certain length or height. WHZ is an indicator of body fat with higher values representing excess weight or obesity. All indicators were expressed as sex and age (in months) specific z-scores, and were analysed as continuous variables.

HFI level (security, mild insecurity, moderate insecurity, severe insecurity) was the key independent variable and was measured with the previously used and extensively validated Brazilian Food Insecurity Scale (EBIA). EBIA represents an adaptation of the US Household Food Security Survey Module (US-HFSSM), developed during the early 90's and first fielded in the 1995 US Current Population Survey. The detailed description of the adaptation and validation of the EBIA scale can be found elsewhere [23,24], but it is important to state that several validity criteria (content, face, predictive and convergent) were met. The EBIA is composed of 15 dichotomous (yes/no) questions that evaluate food insecurity experiences, ranging from the worry or concern that the household may run out of food, to sacrificing the quality of the diet and to restricting the amount of food consumed, and ultimately going for a whole day with little or no food due to economic limitations. Each household is assigned a summative food insecurity score based on the number of affirmative responses to the scale items. Households were classified either as food secure (HFI score = 0), mildly food insecure (score = 1–5), moderately food insecure (score = 6–10), or severely food insecure (score = 11–15).

As in our previously published studies with adolescent [20] and adult [1] women, the other key confounders included in the analyses were: mother's self-reported skin colour/ethnicity (white, black, brown, yellow, indigenous); maternal schooling in years (0–4, 5–8, ≥ 9); place of household (urban, rural); geographical region of the household (North, Northeast, Southeast, South, Midwest); quartiles of family income (in Brazilian *reais*, about US\$ 0,47 per *real* in 2006–07); mothers' smoking habit (yes, no), mother's marital status (single/widowed/divorced, married or cohabiting), number of children 0–60 months in the household and birth order. These covariates were selected based on theoretical and empirical considerations and allows for direct comparison between the findings of the present study with those of previous studies [1,20].

Post-hoc sample size calculations

Since the number of children available for analysis was fixed and based on the number available on the DHS database post-hoc sample size calculation was performed. A sample size of 157 children in each HFI group is able to detect a difference of 0.1 z-

Table 1. Frequency distribution of Household Food Insecurity (HFI) levels according to socio-demographic variables.

| | Food security | | HFI level | | | | | | <i>p</i> ² |
|--|---------------|-------|-----------|-------|----------|-------|--------|-------|-----------------------|
| | n | % | Mild | | Moderate | | Severe | | |
| | | | n | % | n | % | n | % | |
| Geographical region of household | | | | | | | | | <0.001 |
| North | 312 | 39.43 | 193 | 24.15 | 153 | 19.63 | 154 | 16.79 | |
| Northeast | 250 | 42.17 | 203 | 28.55 | 155 | 21.07 | 73 | 8.21 | |
| Southeast | 389 | 55.19 | 196 | 33.36 | 60 | 9.45 | 19 | 2.00 | |
| South | 385 | 67.20 | 154 | 23.48 | 43 | 5.88 | 20 | 3.44 | |
| Midwest | 374 | 59.31 | 198 | 27.08 | 70 | 9.92 | 32 | 3.69 | |
| Place of household | | | | | | | | | 0.277 |
| Urban | 1,163 | 51.16 | 653 | 30.41 | 319 | 12.89 | 184 | 5.54 | |
| Rural | 547 | 52.65 | 291 | 24.66 | 162 | 16.29 | 114 | 6.40 | |
| Family income (quartiles) | | | | | | | | | <0.001 |
| 1 | 337 | 30.99 | 334 | 32.66 | 275 | 24.67 | 218 | 11.68 | |
| 2 | 479 | 44.20 | 321 | 34.99 | 150 | 14.84 | 60 | 5.97 | |
| 3 | 453 | 64.60 | 210 | 29.71 | 46 | 4.80 | 15 | 0.89 | |
| 4 | 441 | 85.15 | 79 | 12.79 | 10 | 1.70 | 5 | 0.36 | |
| Mother's skin color | | | | | | | | | <0.001 |
| White | 709 | 61.58 | 274 | 24.96 | 103 | 8.92 | 62 | 4.54 | |
| Black | 133 | 38.75 | 109 | 37.23 | 65 | 19.42 | 28 | 4.60 | |
| Brown | 796 | 47.89 | 517 | 30.52 | 187 | 15.11 | 179 | 6.48 | |
| Yellow | 51 | 55.60 | 24 | 28.14 | 11 | 11.81 | 4 | 4.45 | |
| Indigenous | 21 | 23.16 | 20 | 34.01 | 15 | 22.53 | 25 | 20.30 | |
| Mother's years of schooling | | | | | | | | | <0.001 |
| ≥9 | 904 | 66.91 | 329 | 24.77 | 83 | 5.98 | 40 | 2.34 | |
| 5–8 | 530 | 42.87 | 367 | 33.77 | 240 | 18.84 | 93 | 4.53 | |
| 0–4 | 276 | 30.98 | 248 | 31.92 | 158 | 21.17 | 165 | 15.92 | |
| Mother's smoking habit | | | | | | | | | 0.001 |
| No | 1,527 | 54.44 | 796 | 27.53 | 383 | 12.62 | 238 | 5.41 | |
| Yes | 183 | 35.13 | 148 | 39.33 | 98 | 18.24 | 60 | 7.30 | |
| Mother's marital status | | | | | | | | | 0.194 |
| Single/window/divorced | 202 | 42.11 | 109 | 34.26 | 72 | 19.07 | 45 | 4.56 | |
| Married or cohabiting | 1,508 | 52.72 | 835 | 28.70 | 409 | 12.73 | 253 | 5.85 | |
| Sex | | | | | | | | | 0.820 |
| Boys | 874 | 50.42 | 476 | 29.89 | 254 | 14.19 | 146 | 5.50 | |
| Girls | 836 | 52.53 | 468 | 28.81 | 227 | 12.74 | 152 | 5.92 | |
| Age (months) | | | | | | | | | 0.568 |
| 0–24 | 1,018 | 52.71 | 550 | 29.01 | 255 | 13.12 | 150 | 5.16 | |
| 25–60 | 692 | 49.61 | 394 | 29.89 | 226 | 14.03 | 148 | 6.47 | |
| Birth order | | | | | | | | | <0.001 |
| 1 | 777 | 58.06 | 392 | 29.27 | 124 | 9.88 | 49 | 2.79 | |
| 2 | 556 | 52.40 | 260 | 28.27 | 128 | 14.79 | 62 | 4.54 | |
| ≥3 | 377 | 36.22 | 292 | 31.01 | 229 | 19.48 | 187 | 13.29 | |
| Number of children 0–60 months living in the household | | | | | | | | | <0.001 |
| 1 | 1,131 | 56.93 | 558 | 28.52 | 202 | 10.51 | 107 | 4.04 | |
| 2 | 479 | 42.24 | 291 | 32.55 | 181 | 17.96 | 115 | 7.25 | |
| ≥3 | 100 | 28.60 | 95 | 23.94 | 98 | 28.63 | 76 | 18.83 | |

Demographic Health Survey, Brazil 2006.¹¹The column "n" presents real sample values and the column "%" presents values expanded to the Brazilian population.²P values refer to the Chi-square test for differences in proportions.

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score between two HFI groups for each anthropometric indicator, working with 80% of power and assuming a standard deviation of 0.4 z-score based on our sample data. Sample size calculations were performed in Stata 12.0 using the *sampsi* command.

Analytical sample bias analyses

Analytical sample bias analyses were conducted using an weighted (which took into account the complex sample design) chi-square test for categorical variables in order to understand if the overall profile of the anthropometric outliers (n = 94) and those excluded due to missing key variables (n = 1,613), differed from those included in the analyses (n = 3,433). Comparisons were performed in regards to HFI level prevalence, geographical region and other key confounders.

Statistical Analyses

We first analysed HFI (mild, moderate and severe food insecurity) prevalence distribution according to key confounding variables. Comparisons were performed using weighted chi-square test for differences in proportions. In a second stage, we examined the mean values (with 95% CI) for the two outcomes (WHZ and BMI-Z), according to HFI. Third, interaction terms between HFI and geographical region, child gender and age on each of the two anthropometric outcomes were tested.

We conducted crude and adjusted linear regression analyses (B coefficients and 95% CI), adjusting for key confounders and having WHZ and BMI-Z as outcomes. Separate independent crude and adjusted linear regression models were estimated. Estimates were weighted and the 95% CI corrected to take into account the complex sampling design by means of *svy* commands in Stata (Stata 12.0 software, StataCorp, Texas, USA). Moderate and severe HFI categories were combined to increase the precision and study power.

Lastly, we performed three further analyses to assure that the analytical approach used has produced unbiased and valid results. First we re-analyzed the data including only one child per household, the youngest child (n = 2849, representing 83% of the original sample), considering that 18% of the households had more than one child. Further we run a model with adjustment for mother’s weight and height and finally, since mother’s schooling or income could act alongside HFI or via HFI on BMI-Z or WHZ we have run models removing these variables.

Results

The likelihood of being excluded as outliers according to the anthropometric z-score values (n = 94) or due to missing information for the key variables (n = 1,619) was higher for children living in food secure households, from the South geographical region, for those living in rural areas and for children of mothers’ with schooling <4 years, of white skin color and single or divorced, when compared to those included (n = 3,433) (**Table S1**).

The weighted overall prevalence of any level of HFI was 48.6% (results not shown in tables). Severe HFI was more prevalent among children living in the North region, living in lower income families, with black or brown skin mothers, with mothers who had 0–4 years of schooling, and living in households with ≥3 children of age (**Table 1**).

BMI-Z mean values were lower for children who lived in moderate to severe food insecurity compared with mild and food-secure households. A similar pattern was observed for WHZ (**Table 2**).

We did not find a significant interaction between HFI and geographical region for BMI-Z (P = 0.119) or WHZ (P = 0.198). Thus, all analyses were conducted for the country as a whole and were not stratified by geographical region (**results not shown in tables**).

HFI was negatively associated with BMI-Z (moderate to severe HFI: $\beta = -0.19$, 95% CI: $-0.35 - -0.03$, P = 0.047), and WHZ (moderate to severe HFI: $\beta = -0.43$, 95% CI: $-0.26 - -0.09$, P = 0.009) in the unadjusted analysis. Estimates lost significance after analysis were adjusted for mothers’ skin color, mothers’ years of schooling, place of household, household income quartiles, mothers’ smoking habit, mothers’ marital status, number of children 0–60 months in the household, and birth order (**Table 3**).

Further analyses including only one child per household revealed that key results, specifically the unadjusted and adjusted coefficients for BMI-Z and WHZ barely changed and most importantly, kept the direction of the association. The model adjusted for mother’s weight and height did not affect the main results. The same happened when mother’s schooling and income were removed (data not shown, available on request).

Discussion

The main finding of this study is an unadjusted relationship between HFI and children’s BMI-Z for age or WHZ. However, this association lost significance after adjusting for known confounders including mother’s skin color, mother’s years of schooling, place of household, quartiles of family income, mother’s smoking habit, mother’s marital status, number of children 0–60 months in the household and birth order. It is important to note however that overweight and obesity prevalence’s in Brazil are still low among children 0–60 months. These findings are in partial agreement with a previous analysis [16] and add to it substantially by carefully examining the relationships within Brazilian geographical regions, by including BMI-Z and WHZ as outcomes of interest, and by adjusting for the same confounders included in a previously published analyses examining the HFI-obesity relationship among adolescent [20] and adult [1] Brazilian women. However, the current findings are not in full agreement with studies conducted in other low- and middle-income countries such as Colombia [14,25–26], Pakistan [27], Jamaica [17], Korea [18] and Mexico [19]. This great heterogeneity in results may be explained by the inclusion of different child age groups across studies, the use of different scales to measure HFI and different approaches to analyze food insecurity data (e.g., continuous vs. categorical), and/or the possibility that the relationships between

Table 2. Body mass index for age (BMI-Z) and weight for age (WHZ) means according to the Household Food Insecurity (HFI) level.

| | BMI-Z | | | WHZ | | |
|---------------------------------|-------|------|-----------|-------|------|-----------|
| | n | Mean | 95% CI | n | Mean | 95% CI |
| HFI level | | | | | | |
| Food security | 1,710 | 0.52 | 0.44–0.61 | 1,170 | 0.48 | 0.40–0.57 |
| Mild food insecurity | 944 | 0.50 | 0.39–0.61 | 944 | 0.45 | 0.34–0.56 |
| Moderate+severe food insecurity | 779 | 0.33 | 0.20–0.46 | 779 | 0.22 | 0.08–0.37 |
| Total | 3,433 | 0.48 | 0.42–0.54 | 3,433 | 0.42 | 0.36–0.49 |

Demographic Health Survey, Brazil 2006.[†]
[†]Estimates were weighted and 95% CI were corrected to take into account the complex sampling design.
CI: confidence interval.
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Table 3. Crude and adjusted linear regression models for the effect of the Household Food Insecurity (HFI) level on body mass index for age (BMI-Z) and weight for height (WHZ).

| BMI-Z | | | | | | |
|----------------------------|----------|---------------|-------------|-------|-------------------------------|------------|
| | <i>n</i> | crude β | 95% CI | p^2 | adjusted ³ β | 95% CI |
| Household Food Insecurity | 3,433 | | | 0.047 | | |
| Security | 1,710 | reference | - | | reference | - |
| Mild insecurity | 944 | -0.02 | -0.16-0.11 | | 0.01 | -0.12-0.14 |
| Moderate+severe insecurity | 779 | -0.19 | -0.35--0.03 | | -0.09 | -0.27-0.08 |
| | | | | | | p^2 |
| | | | | | | 0.501 |
| WHZ | | | | | | |
| | <i>n</i> | crude β | 95% CI | p^2 | adjusted ³ β | 95% CI |
| Household Food Insecurity | 3,433 | | | 0.009 | | |
| Security | 1,710 | reference | - | | reference | - |
| Mild insecurity | 944 | -0.03 | -0.17-0.11 | | 0.02 | -0.11-0.15 |
| Moderate+severe insecurity | 779 | -0.26 | -0.42--0.09 | | -0.13 | -0.32-0.05 |
| | | | | | | p^2 |
| | | | | | | 0.292 |

Demographic Health Survey, Brazil 2006.¹

¹Estimates were weighted and 95% CI were corrected to take into account the complex sampling design;

²*p* values refers to the Wald test for multiple comparisons;

³Adjusted for geographical region of household, place of household, household income quartiles, mothers' skin color, mothers' years of schooling, mothers' smoking habit, mothers' marital status, child sex, age, order of birth and number of children 0-60 months of age living in the household.

CI: confidence interval.

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HFI and child anthropometric indicators is highly context specific. Longitudinal designs are needed to disentangle these complex relationships. These studies will need to take into account potential effect modification by geographical region, gender, age and the child birth-weight [15,17].

Following a life-course approach, we have previously documented in Brazil, that HFI is associated with obesity risk among adult and with excessive weight among adolescent women investigated by the 2006 DHS [1,20]. Thus, an important question becomes: why the association between HFI and higher BMI-Z or WHZ mean values was not found among any of the sub-groups of Brazilian children 0–60 months of age that we examined?

A potential explanation for the lack of association between HFI and higher BMI-Z means values in Brazilian children studied by the DHS relies on the different velocities and stages of the nutrition transition experimented by diverse socio-economic status sub-groups and individuals at different periods of the life-cycle, during the last decades. It has already been documented earlier that the dynamics in regards to the occurrence of a shift from under-nutrition to obesity tends to follow a pattern that usually occurs first in adults, followed by adolescents, and only later in time by children [28]. This nutrition transition pattern has been extensively documented for the Brazilian population having as reference some of the most important nationally representative surveys from the last 20 years [28,29].

On the other hand, the deleterious effect of the HFI represented mainly by an increased prevalence of obesity or higher BMI-Z mean population values, may constitute a feature of a very advanced stage of the nutrition transition, and thus has been observed only in adults and adolescents, the first groups to achieve the new pattern, but not in children 0–60 months, the group that usually suffers the nutrition transition effects at a later stage. We can conclude that the nutrition transition happening in Brazil may still be having less of an impact among young children when compared to older children, youth and adults living in food insecure households, when the country as a whole is considered. A cross-sectional study conducted in Pelotas, southern of Brazil, thus a more developed region, has shown that excessive weight for children under 5 years, in families living in food insecurity [30]. These results support our hypothesis. Alternatively the high nutrient demands that young children have per unit of body weight may also partially explain why the obesity risk/higher BMI-Z mean values, among children as a result of exposure to HFI is differentially affected in relationship to adult women. A recent study conducted on a vulnerable urban area of Rio de Janeiro, Brazil, has shown that children between 6 and 30 months living in food insecure households were more likely to have lower quality diets, characterized by highly energy dense foods and low protein content [31]. It is possible that the relatively high caloric

requirements per unit of body weight that children of this age have, may “protect” them against becoming obese, even if they are exposed to cheap energy dense diets that may result from HFI. It is important to monitor if and how HFI coping mechanisms differ (e.g. dietary adaptation) across age groups in the context of the nutrition transition.

The cross-sectional nature of the study, as well as the lack of information on coping mechanisms associated with diverse levels of HFI should be considered as a limitation. The evidence for a relationship between food insecurity and nutritional status is still inconsistent, as almost all studies are cross-sectional, precluding the possibility of establishing the temporality of events. In this regard, it is essential to conduct prospective studies that collect data on HFI status across time, as well as information as to how children and adults within the household cope with HFI changes. Special attention needs to be paid to dietary intake and physical activity adaptations. Ultimately the goal is to better understand how these coping behaviours influence body fat accumulation at different stages of the life cycle and across countries undergoing different phases of the nutrition transition. In addition it is important to acknowledge that complex surveys as DHS often times have substantial missing data in key self-reported confounders such as birth weight and digit preference biases in indicators such as breastfeeding duration. Our study is prone to some sort of selection bias, because there were statistically significant differences when comparing included vs. excluded subjects due to missing values. The likelihood of being excluded was higher for children with food security, from the South geographical region, for children of mothers’ with schooling <4 years and single or divorced. Thus, the results need to be interpreted with caution.

This study detected that there was no association between HFI and BMI-Z or WHZ values in Brazilian children 0–60 months of age. These results together with empirical evidence from Brazil and other countries suggests that the nutrition transition phase at which countries are needs to be taken into account when interpreting findings from studies examining the relationship between HFI and obesity risk across the life cycle.

Supporting Information

Table S1 Frequency distribution of main investigated variables between those who were included in the study and those who were not. Demographic Health Survey, Brazil 2006.
(DOC)

Author Contributions

Conceived and designed the experiments: GK RP-E. Analyzed the data: GK MMS AAMS GV-M. Wrote the paper: GK MMS RP-E AAMS GV-M.

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