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RESEARCH ARTICLE

Obstetric fistula repair failure and its associated factors among women who underwent repair in sub-Saharan Africa. A systematic review and meta-analysis

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Abstract

Background

Obstetric fistula repair failure can result in increased depression, social isolation, financial burden for the woman, and fistula care programs. However, there is limited, comprehensive evidence on obstetric fistula repair failure in Sub-Saharan African countries. This systematic review and meta-analysis aimed to determine the pooled prevalence of obstetric fistula repair failure and associated factors among women who underwent surgical repair in Sub-Saharan African countries.

Methods

To identify potential articles, a systematic search was done utilizing online databases (PubMed, Hinari, and Google Scholar). The Preferred Reporting Items for Systematic Review and Meta-Analysis Statement (PRISMA) guideline was used to report the review's findings. I² test statistics were employed to examine study heterogeneity. A random-effects model was used to assess the pooled prevalence of obstetric fistula repair failure, and the association was determined using the log odds ratio. Publication bias was investigated using the funnel plot and Egger's statistical test at the 5% level of significance. Meta-regression and subgroup analysis were done to identify potential sources of heterogeneity. The data were analyzed using STATA version 17 statistical software.

Results

A total of 24 articles with 9866 study participants from 13 Sub-Saharan African countries were included in this meta-analysis. The pooled prevalence of obstetric fistula repair failure in sub-Saharan Africa was 24.92% [95% CI: 20.34–29.50%]. The sub-group analysis by country revealed that the highest prevalence was in Angola (58%, 95% CI: 53.20–62.80%)

Abbreviations: BMI, Body Mass Index; CoCoPop, Condition, Context, Population; CI, Confidence Interval; IPS, Infection Prevention Strategies; NOS, Newcastle Ottawa Scale; OR, Odds Ratio; SDG, Sustainable Development Goals; UN, United Nations; PROSPERO, Prospective Register of Systematic Reviews; PRISMA, Preferred Reporting Items for Systematic Reviews and Meta-Analyses; SSA, Sub-Saharan Africa; SDG, Sustainable Development Goals; WHO, World Health Organization. and the lowest in Rwanda (13.9, 95% CI: 9.79–18.01%). Total urethral damage [OR = 3.50, 95% CI: 2.09, 4.91], large fistula [OR = 3.09, 95% CI: (2.00, 4.10)], duration of labor [OR = 0.45, 95% CI: 0.27, 0.76], and previous fistula repair [OR = 2.70, 95% CI: 1.94, 3.45] were factors associated with obstetric fistula repair failure.

Conclusion

Women who received surgical treatment for obstetric fistulas in Sub-Saharan African countries experienced more repair failures than the WHO standards. Obstetric fistula repair failure was affected by urethral damage, fistula size, duration of labor, types of fistula, and history of previous repairs. Therefore, we suggest policy measures specific to each country to provide special attention to the prevention of all risk factors, including poor nutrition, multiparty, obstructed labor, and maternal age, which can result in conditions like large fistulas, urethral damage, and repeat repair, in order to reduce obstetric fistula repair failure.

Introduction

An obstetric fistula is defined by the World Health Organization (WHO) as "an abnormal opening between a woman's vagina and bladder and/or rectum through which her urine and/ or feces continuously leak" [1]. These openings can be vesico-uterine (between the bladder and uterus), rectovaginal (between the rectum or colon and the vagina), ureterovaginal (between the ureter and the vagina), and/or vesicovaginal (between the anterior vaginal wall and posterior bladder) [2]. Obstetric fistula is frequently the result of obstructed labor, which happens when the baby's head gets caught in the mother's pelvis, obstructing blood flow to the neighboring tissues and leading to tissue necrosis, which can result in a fistula [2, 3]. In developing countries, aside from obstructed labor, other factors that contributed to the occurrence of obstetric fistulas included a lack of skilled birth attendants, poor health-seeking behavior, a poor referral system, poor transportation, age and physical maturity, iatrogenic surgical damage, sexual violence, poverty, a lack of awareness, and not spacing out pregnancies [4–6].

The WHO estimates that obstetric fistulas affect between 50 000 and 100 000 women annually. It is predicted that 30 000–130 000 new cases of obstetric fistulas develop each year in Africa alone [7], with over 2 million young women in Asia and Sub-Saharan Africa (SSA) not obtaining treatment for their condition [8]. In nearly 60 countries, it is challenging. Additionally, obstetric fistula can be found in all developing nations, including those in Africa, the Middle East, and Asia. It is most common in the region of SSA known as the fistula belt, which spans the northern half of the continent from Mauritania to Eritrea [9].

Obstetric fistulas have major physical, social, and psychological effects if early, efficient medical care is not received. It causes fecal, urinary, or both types of incontinence [10, 11]. The vulva and thighs may get damaged as a result of ongoing urine and feces leakage [2]. In addition, women with obstetric fistulae are more likely to experience social isolation and marginalization [12, 13], high rates of divorce or separation [14, 15], lack of sexual desire [13, 14], loss of fertility and amenorrhea [16], depression [15, 17], associated comorbidities, reduced self-esteem, and income loss [10, 11, 18].

Obstetric fistula has been a neglected public health and human rights issue despite the fact that it is currently almost entirely preventable and is a sign that social and health systems are failing to protect the health and human rights of the poorest and most vulnerable women and

girls. It also persists as a reminder of egregious injustices, a symbol of global inequality, and a sign that these women's and girls' rights are not being protected [19, 20].

Obstetric fistula repair surgery is the only proven treatment for the condition, yet results can vary widely based on the patient's location and the healthcare setting. Repairing an obstetric fistula is essential for a woman's general health and well-being. The treatment of an obstetric fistula necessitates a precise diagnosis, preoperative care, quick repair, application of fundamental surgical concepts, postoperative care, and follow-up. In addition to causing pain, both physical and psychological, failing to repair a fistula costs money and jeopardizes efforts to end this treatable and preventable illness. The United Nations (UN) has also established a goal to eliminate obstetric fistulas by 2030; this can be accomplished as more women have access to competent birth attendants and as access to expert obstetric fistula repair has improved [20–22]. The WHO established a goal of fewer than 15% for failed fistula closure following repair and less than 10% for incontinence after successful fistula closure as the best range of repair outcomes for assessing the quality of services provided to patients [1].

According to studies done in the SSA, failure rates for obstetric fistula repair vary significantly depending on the context, with failure rates ranging from 11% to 58% [23–28]. These studies have outlined the rates of obstetric fistula repair failure as well as a variety of risk factors. However, the majority of these investigations are constrained by the small sample sizes, varied repair failure definitions, varied quality, and the fact that the majority of the studies were carried out at a single center in a single country, which limits the ability of these studies to generalize beyond a very specific patient population and setting. Additionally, to the best of our knowledge, there is also a limited, comprehensive analysis of the available evidence on obstetric fistula repair failure and the associated factors among women who underwent surgical repair in SSA.

Therefore, the purpose of this systematic review and meta-analysis was to determine the pooled prevalence of obstetric fistula repair failure and the associated factors among women who underwent surgical repair in SSA. This study offers a comprehensive review of the data available from existing sources to produce evidence-based information for obstetric care providers, policy planners, ministries of health, and other relevant stakeholders in SSA regarding the prevalence of obstetric fistula repair failures and associated risk factors. It also serves as a foundation for future prospective studies in the field and aims to improve the effectiveness, quality, and outcomes of fistula repair interventions in SSA.

Materials and methods

Protocol registration and reporting

The protocol for this systematic review and meta-analysis was developed and registered on the PROSPERO (prospective register of systematic reviews) (registration number: CRD42023409402). The Preferred Reporting Items for Systematic Review and Meta-Analysis Statement (PRISMA) guideline was used to report the review's findings [29] (S1 File).

Study design, setting, and search strategy/data source

To determine the pooled prevalence of obstetric fistula repair failure and associated factors among women who had repair, a systematic review and meta-analysis were carried out in SSA. The region of Africa known as SSA, which includes West Africa, Southern Africa, East Africa, and Central Africa, is situated south of the Sahara.

The studies were retrieved online using MEDLINE/PubMed, Google Scholar, Hinari, and Gray (unpublished) literature, academic archives, and manually searching references from a list of included. The search was conducted using the following keywords, either separately or

in combination: prevalence, magnitude, repair failure, unsuccessful closure, successful repair with incontinence, associated factors, predictors, determinants, contributing factors, obstetric fistula, vesicovaginal fistula, rectovaginal fistula, women, female, SSA. The following keywords were utilized to obtain articles from PubMed: (Epidemiology) OR (Prevalence) OR (Magnitude) AND (Obstetric Fistula) OR (Vesico vaginal Fistula) OR (Recto vaginal Fistula) OR (Ureterovaginal Fistula) OR (Vesico uterine Fistula) AND (Unsuccessful fistula closure) OR (Incontinence following a successful closure) OR (Repair failure) AND (Associated factors) OR (Determinants) OR (Predictors AND) AND (Women who underwent obstetric fistula repair) AND (Sub-Saharan Africa) OR (SSA). Three authors (HEH, ZA, and MA) searched for articles during the months of May 1 and July 1, 2023, with no restrictions on the date of publication. Using the EndNote Version X7 reference manager, articles found through electronic searches were exported, organized, and duplicate results were removed.

Study selection and eligibility criteria

Before retrieving the full-text articles, two researchers (HEH and BGD) independently reviewed the titles and abstracts of the chosen studies. The full-text articles were further screened using predetermined inclusion criteria. During a consensus meeting with the other reviewers (MA and ZA), disagreements regarding the final selection of studies to be included in the systematic review and meta-analysis were discussed and solved. To establish inclusion and exclusion criteria for prevalence studies, we employed the CoCoPop (Condition, Context, and Population) approach.

Inclusion criteria. *Participants.* Participants were women who had undergone obstetric fistula surgery or treatment.

Study settings. Studies carried out in SSA countries in community or institutional settings.

Study design. All observational studies (i.e., cross-sectional, case-control, and cohort) reporting the magnitude of obstetric fistula repair failure and its associated factors were eligible for this review.

Publication type. Both published (journal articles) and unpublished (master's theses and dissertations) articles without restriction of date of publication were included.

Language. The review included only English-language studies.

Exclusion criteria. Articles addressing rape-related fistulas or those caused by non-obstetric causes (such as hysterectomy)were excluded from the study. Editorials, letters, reviews, comments, interventional studies, articles without abstracts and whose full data were not accessible despite requests from the authors, duplicate studies, and articles with poor methodological quality were excluded. Studies that only showed qualitative data on obstetric fistula repair failure in women who had repairs were also excluded.

Measuring outcome variables

There are two main outcomes of this study. The primary outcome variable for this study, obstetric fistula repair failure, is the proportion of women with failed fistula closure and/or incontinence after successful fistula closure at 21 days following surgery. In order to determine the prevalence of obstetric fistula failure, the total number of study participants (including the sample) was divided by the number of women who experienced failed fistula closure and/or incontinence following successful fistula closure and then multiplied by 100. Identifying the determinants of obstetric fistula failure in women who have had repairs was the second outcome of this study. In order to establish the association between obstetric fistula repair failure and factors for the second outcome, we calculated the log odds ratio. Based on binary results from the primary studies, the odds ratio for significant factors was estimated.

Data extraction

Using a pre-piloted data extraction format prepared in a Microsoft Excel spreadsheet, two authors (HEH and MA) independently extracted all essential data. All discrepancies between the two writers were settled by conversation and agreement at the time of data collection. The principal author of the original study was contacted if more details or clarifications were required. The primary author name, publication year, countries where the study was done, study setting, sample size, study design, response rate, and prevalence of obstetric fistula repair failure were all included in the data extraction form for the first outcome. For the secondary outcome (associated factors), data were retrieved in the form of two by two tables, and the log odds ratio for each factor was then calculated based on the findings of the original research.

Assessment of the quality of the individual studies

The quality of each article chosen for review was independently evaluated by two reviewers (HEH and ZA). The methodological quality of possible studies was evaluated using the Newcastle-Ottawa scale (NOS) [30], a tool for evaluating the quality of observational studies in systematic reviews and meta-analyses, in order to determine the likelihood of bias within the included research. Study group selection was evaluated based on sample representativeness, exposure determination, sample size, and non-response rate; comparability was evaluated based on subject comparability; and outcome was evaluated based on an assessment of the outcome and a statistical test for cross-sectional studies. The NOS ratings are used to evaluate the articles' quality and eligibility in four main categories. Category I: Selection (five points) because: ascertainment of exposure (2 points), representativeness of the sample (1 point), sampling methodology (1 point), and response rate (1 point); category II: comparability (2 points): confounding controlled (data or results adjusted for pertinent predictors, risk variables, and confounding factors (2 points); category III: outcome (3 points): assessment of outcome (2 points) and statistical tests (1 point). The next step was to calculate the quality score for each study, which for cross-sectional studies ranged from zero to ten. Finally, the meta-analysis only included high-quality publications that had a score of at least 6 out of 10. Any disagreements that arose between the reviewers were resolved through conversation or with another reviewer (BGD and MA).

Data processing and statistical analysis

For the meta-analysis, the retrieved data were exported into STATA version 17. In order to account for the observed variability, a random-effects (DerSimonian and Laird) method was used in a meta-analysis to determine the pooled prevalence of obstetric fistula repair failure among women who had repair. An independent analysis was also done on the effects of particular determining factors on obstetric fistula repair failure. A forest plot was created and used to display the pooled effect size (prevalence and odds ratio (OR)), along with a 95% confidence interval (CI). I² statistics and Cochran's Q test were used to assess the heterogeneity between studies. The percentage of the total variation in the study estimate that is related to heterogeneity is measured using I². I² values vary from 0 to 100%, and those greater than 75% indicate significant study heterogeneity. A statistically significant heterogeneity was defined as having a p-value of less than 0.1. A subgroup analysis was performed with regard to the country, year of publication, and sample size in order to reduce the variance of estimated points between primary studies. Additionally, a univariate meta-regression analysis was conducted using the sample size, year of publication, and response rate as covariates.

A funnel plot was used to analyze publication bias graphically. The funnel plot's asymmetry is an indication of possible publication bias. A p-value of less than 0.05 was used to determine

the presence of significant publication bias using Egger's test, indicating the existence of a small study effect [31]. Then, a trim and fill analysis was conducted to deal with publication bias. Moreover, sensitivity analysis was carried out to see whether individual studies had an impact on the pooled effect size. To ascertain the relationship between related factors and obstetric fistula repair failure, a log odds ratio with a 95% confidence interval (CI) was utilized. A forest plot and the odds ratio (OR) with its 95% CI were used to present the results of the meta-analysis.

Result

Study selection and identification

A total of 16,504 articles were identified using an electronic database search; the manual search did not add additional studies. Of these, Endnote citation manager software version X7 deleted 14850 duplicate articles, and 1528 studies were excluded by reviewing the titles and abstracts of the remaining articles. Out of the 126 articles that were left, 102 were disregarded because they did not include the desired outcome, had inconsistent result reports, had irrelevant target groups, or had duplicated information. Finally, a total of 24 articles that satisfied the criteria for inclusion, representing 13 SSA countries, were included to the study (Fig 1).

Characteristics and quality of the included articles

In total, 24 studies that were carried out in SSA nations were included in this meta-analysis. Regarding the year of publication, all research was published between 2007 and 2022. Thirteen articles were published during and before 2016, and eleven of the studies were published after 2016, all of which were conducted in hospitals (100%), and which took place in 13 SSA countries. In total, 9866 women who had obstetric fistulas repaired participated in the 24 studies that were included and carried out in SSA. The mean number of samples was 411, while the minimum and maximum sample sizes were 32 and 1303, respectively. In terms of study design, there were nine cohort studies, six cross-sectional studies, eight retrospective chart reviews, and one case-control study were among the included articles.

Out of the 24 articles, 5 were from Ethiopia [26, 32–35], 4 were from the Democratic Republic of the Congo [25, 36–38], 2 were from Nigeria [23, 39], 1 was from Niger [40], 2 were from Uganda [24, 41], 2 were from Zambia [16, 42], 2 were from Kenya [27, 43], 1 was from Guinea [44], 1 was from Angola [28], 1 was from Benin [45], 1 was from Rwanda [46], 1 was from Tanzania [47], and 1 was from Burkina Faso [48]. All of the included studies were published.

According to the articles that were reviewed, the prevalence of obstetric fistula repair failure ranged from the lowest 11% in Ethiopian [26], Nigerian [23], and Ugandan [24] studies to the highest 58% in an Angolan study [28].

During our quality assessment, we verified that all included studies had reliable methodological quality (NOS scores vary from 7 to 10 out of a possible 10-point total). It determined that there was moderate to almost complete agreement among investigators regarding the level of bias for the studies that were included in the final analysis (Table 1).

Prevalence of obstetric fistula repair failure in SSA

In order to assess the Der Simonian and Laird overall effect, a random-effects model was utilized because there was significant heterogeneity among the included studies ($I^2 = 97.16\%$, p-value < 0.001). Therefore, in SSA, the pooled prevalence of obstetric fistula repair failure among women who underwent repair was 24.92% (95% CI: 20.34–29.50%) (Fig 2).



Fig 1. PRISMA flow diagram demonstrating the literature search and screening process.

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Subgroup analysis by country, sample size, and publication year was carried out to identify the possible source of heterogeneity between studies.

On country-wise subgroup analysis, the country where the study was conducted was one of the major sources of significant heterogeneity. The highest prevalence of obstetric fistula repair

| Authors name | Publication year | Country | Centers | Sample size | Outcome | Response rate (%) | Prevalence (%) | Quality Score |
|--|---------------------|-----------------|---|----------------|---------|----------------------|-------------------|------------------|
| Ahmed ZA et al. [23] | 2013 | Nigeria | Laure fistula center | 343 | 32 | 83.9 | 11.1 | 7 |
| Aynie AA et al. [32] | 2019 | Ethiopia | Bahir Dar, Hamlin Fistula Center | 385 | 136 | 100 | 35.3 | 10 |
| Bello OM et al. [39] | 2011 | Nigeria | Ibadan, tertiary institution | 155 | 43 | 77 | 36.1 | 8 |
| Bernard L et al. [28] | 2019 | Angola | Centro Evangélico de Medicina do Lubango (CEML) hospital | 407 | 234 | 100 | 58 | 8 |
| Delamou A et al. [<u>44]</u> | 2016 | Guinea | Three Engender health supported repair hospitals | 785 | 109 | 96 | 14.5 | 9 |
| G. Egziabher T et al. [<u>46</u>] | 2015 | Rwanda | Public tertiary hospital | 272 | 38 | 100 | 13.9 | 9 |
| Goh GTW et al. [<u>33]</u> | 2008 | Ethiopia | Addis Ababa hospital and Bahirdar Hamlin fistula unit | 1024 | 256 | 96.4 | 25.9 | 8 |
| Hawkins L et al. [43] | 2013 | Kenya | Three medical centers in western Kenya | 556 | 87 | 86.7 | 18 | 8 |
| Holme A et al. [16] | 2007 | Zambia | Zambia's primary fistula repair center, Monze Mission Hospital. | 259 | 69 | 98.4 | 27.1 | 7 |
| Holt L et al. [24] | 2021 | Uganda | Fistula Hospital in Uganda | 546 | 59 | 99.1 | 11 | 9 |
| Imakando MM et al. [42] | 2022 | Zambia | Managedata Teaching Hospital | 32 | 3 | 56 | 17 | 7 |
| Kayondo M et al. [<u>41</u>] | 2011 | Uganda | Mbarara Regional Referral Hospital in western Uganda | 77 | 17 | 100 | 22.1 | 8 |
| Loposso M et al. [38] | 2016 | Dr.Congo | Saint Luc Hospital Kisantu | 166 | 47 | 100 | 28.3 | 7 |
| Mafu MM et al. [25] | 2022 | Dr.Congo | Saint Joseph Hospital in Kinshasa, Panzi Hospital in Bukavu, South Kivu and HEAL Africa Hospital in Goma, North Kivu. | 895 | 106 | 100 | 11.8 | 9 |
| Magali KT et al. [47] | 2016 | Tanzania | Bugando Medical Centre, Mwanza | 159 | 57 | 83 | 42.9 | 7 |
| Mubikayi L et al. [37] | 2017 | Dr.Congo | Four fistula centers, Hospital | 483 | 72 | 100 | 14.9 | 7 |
| Mwangi HR et al. [27] | 2018 | Kenya | Gynocare Women's and Fistula Hospital | 357 | 119 | 88 | 37.9 | 8 |
| Nardos R et al. [26] | 2009 | Ethiopia | Addis Ababa Hamlin Fistula Hospital | 1303 | 115 | 81 | 11 | 8 |
| Nsambi J et al. [<u>36]</u> | 2019 | Dr.Congo | Facilities in the south eastern part of the province of the Upper Katanga | 384 | 66 | 100 | 17.1 | 9 |
| Ouedraogo I et al. [<u>40]</u> | 2018 | Niger | Danja Fistula Center | 384 | 177 | 100 | 46 | 8 |
| Sambo BT et al. [45] | 2016 | Benin | Hospital of Bembereke | 82 | 22 | 100 | 26.8 | 8 |
| Sori DA et al. [34] | 2016 | Ethiopia | Jimma University teaching Hospital | 200 | 26 | 84 | 15.5 | 7 |
| Tadesse S et al. [35] | 2022 | Ethiopia | Yirgalem Hamlin fistula center | 562 | 162 | 100 | 28.8 | 10 |
| Traore TM et al. [48] | 2023 | Burkina Faso | regional teaching hospital | 50 | 16 | 100 | 32 | 8 |

| Table 1. | Characteristics o | f the included | studies for | the systematic | review and meta | -analysis. |
|----------|-------------------|----------------|-------------|----------------|-----------------|------------|
|----------|-------------------|----------------|-------------|----------------|-----------------|------------|

failure was observed among studies conducted in Angola (58.00%, 95% CI: 53.20–62.80%) and the lowest was observed in Rwanda (13.90%, 95% CI: 9.79–18.01%) (Fig 3).

A subgroup analysis was carried out based on the sample size (mean) and publication year (mean). Accordingly, it was found that there was a 28.16% prevalence of obstetric fistula repair

| Study | | | | | witl | prev h 95% CI | Weight (%) |
|--|---|----|----------|----|-------------|------------------|---------------|
| Ahmed ZA et al | - | - | | | 11.10 [| 7.78, 14.42] | 4.38 |
| Aynie AA et al | | | | | 35.30 [3 | 30.53, 40.07] | 4.27 |
| Bello OM et al | | | _ | | 36.10 [2 | 28.54, 43.66] | 3.99 |
| Bernard L et al | | | | | - 58.00 [: | 53.20, 62.80] | 4.27 |
| Delamou A et al | | - | | | 14.50 [| 12.04, 16.96] | 4.42 |
| G. Egziabher T et al | - | - | | | 13.90 [| 9.79, 18.01] | 4.32 |
| Goh GTW et al | | | | | 25.90 [2 | 23.22, 28.58] | 4.41 |
| Hawkins L et al | | - | | | 18.00 [| 14.81, 21.19] | 4.38 |
| Holme A et al | | _ | ⊢ | | 27.10 [2 | 21.69, 32.51] | 4.22 |
| Holt L et al | - | · | | | 11.00 [| 8.38, 13.62] | 4.42 |
| Imakando MM et al | | - | - | | 17.00 [| 3.99, 30.01] | 3.29 |
| Kayondo M et al | | | _ | | 22.10 [| 12.83, 31.37] | 3.79 |
| Loposso M et al | | _ | — | | 28.30 [2 | 21.45, 35.15] | 4.07 |
| Mafu MM et al | 1 | | | | 11.80 [| 9.69, 13.91] | 4.44 |
| Magali KT et al | | | _ | _ | 42.90 [3 | 35.21, 50.59] | 3.98 |
| Mubikayi L et al | | - | | | 14.90 [| 11.72, 18.08] | 4.39 |
| Mwangi HR et al | | | | | 37.90 [3 | 32.87, 42.93] | 4.25 |
| Nardos R et al | | | | | 11.00 [| 9.30, 12.70] | 4.45 |
| Nsambi J et al | | - | | | 17.10 [| 13.33, 20.87] | 4.35 |
| Ouedraogo I et al | | | - | F | 46.00 [4 | 41.02, 50.98] | 4.25 |
| Sambo BT et al | | | <u> </u> | | 26.80 [| 17.21, 36.39] | 3.75 |
| Sori DA et al | | - | | | 15.50 [| 10.48, 20.52] | 4.25 |
| Tadesse S et al | | - | - | | 28.80 [2 | 25.06, 32.54] | 4.35 |
| Traore TM et al | | | | | 32.00 [| 19.07, 44.93] | 3.30 |
| Overall | | • | • | | 24.92 [2 | 20.34, 29.50] | |
| Heterogeneity: $\tau^2 = 122.00$, $I^2 = 97.16\%$, $H^2 = 35.26$ | | | | | | | |
| Test of $\theta_i = \theta_j$: Q(23) = 810.99, p = 0.00 | | | | | | | |
| Test of $\theta = 0$: $z = 10.66$, $p = 0.00$ | | | | | | | |
| | 0 | 20 | 40 | 60 | | | |

Random-effects DerSimonian-Laird model

Fig 2. Forest plot showing the pooled prevalence of obstetric fistula repair failure in Sub-Saharan Africa. Subgroup analysis on the prevalence of obstetric fistula repair failure.

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failure in studies that were finished after 2016 and a 28.82% prevalence in studies with a sample size of less than or equal to 411. In both cases, there was significantly higher heterogeneity among studies (Table 2).

Publication bias

Egger's tests were used to determine publication bias in the studies that were included in the meta-analysis. According to Egger's test, there was evidence of publication bias (P-value ≤ 0.001). The nonparametric rank correlation (Begg) test also revealed there is

| Study | prev with 95% CI | Weig (%) |
|---|-------------------------|-------------|
| Angola | | |
| Bernard L et al | | 4.27 |
| Heterogeneity: $\tau^2 = 0.00$, $I^2 = .\%$, $H^2 = .$ | 58.00 [53.20, 62.80] | |
| Test of $\theta_i = \theta_j$: Q(0) = -0.00, p = . | • | |
| Benin | | |
| Sambo BT et al | 26.80 [17.21, 36.39] | 3.75 |
| Heterogeneity: $\tau^2 = 0.00$, $I^2 = .\%$, $H^2 = .$ | 26.80 [17.21, 36.39] | |
| Test of $\theta_i = \theta_j$: Q(0) = -0.00, p = . | | |
| Burkina Faso | | |
| Traore TM et al | 32.00 [19.07, 44.93] | 3.30 |
| Heterogeneity: $\tau^2 = 0.00$, $I^2 = .\%$, $H^2 = .$ | 32.00 [19.07, 44.93] | |
| Test of $\theta_i = \theta_j$: Q(0) = -0.00, p = . | | |
| Dr.Congo | | |
| Loposso M et al — | - 28.30 [21.45, 35.15] | 4.07 |
| Mafu MM et al | 11.80 [9.69, 13.91] | 4.44 |
| Mubikayi L et al - | 14.90 [11.72, 18.08] | 4.35 |
| Nsambi J et al 🕂 📲 | 17.10 [13.33, 20.87] | 4.35 |
| Heterogeneity: $\tau^2 = 20.92$, $I^2 = 87.33\%$, $H^2 = 7.89$ | 17.13 [12.21, 22.05] | |
| Test of $\theta_i = \theta_j$: Q(3) = 23.67, p = 0.00 | | |
| Ethiopia | | |
| Aynie AA et al - | 35.30 [30.53, 40.07] | 4.27 |
| Goh GTW et al | 25.90 [23.22, 28.58] | 4.41 |
| Nardos R et al | 11.00 [9.30, 12.70] | 4.45 |
| Sori DA et al | 15.50 [10.48, 20.52] | 4.25 |
| Tadesse S et al | 28.80 [25.06, 32.54] | 4.35 |
| Heterogeneity: $\tau^2 = 113.09$, $I^2 = 97.80\%$, $H^2 = 45.46$ | 23.24 [13.76, 32.71] | |
| Test of $\theta_i = \theta_j$: Q(4) = 181.83, p = 0.00 | | |
| Guinea | | |
| Delamou A et al 📲 | 14.50 [12.04, 16.96] | 4.42 |
| Heterogeneity: $\tau^2 = 0.00$, $I^2 = .\%$, $H^2 = .$ | 14.50 [12.04, 16.96] | |
| Test of $\theta_i = \theta_j$: Q(0) = -0.00, p = . | | |
| Kenya | | |
| Hawkins L et al | 18.00 [14.81, 21.19] | 4.38 |
| Mwangi HR et al | 37.90 [32.87, 42.93] | 4.25 |
| Heterogeneity: $\tau^2 = 193.38$, $I^2 = 97.66\%$, $H^2 = 42.82$ Test of $\theta_i = \theta_j$: $Q(1) = 42.82$, $p = 0.00$ | 27.85 [8.35, 47.35] | |
| Niger | | |
| Ouedraogo I et al | 46.00 [41.02, 50.98] | 4.25 |
| Heterogeneity: $\tau^2 = 0.00$, $I^2 = .\%$, $H^2 = .$ | 46.00 [41.02, 50.98] | |
| Test of $\theta_i = \theta_j$: Q(0) = -0.00, p = . | | |
| Nigeria | | |
| Ahmed ZA et al | 11.10 [7.78, 14.42] | 4.38 |
| Bello OM et al | 36.10 [28.54, 43.66] | 3.99 |
| Heterogeneity: $\tau = 303.62$, $\Gamma = 97.16\%$, $H = 35.19$ Test of $\theta_i = \theta_i$: $Q(1) = 35.19$, $p = 0.00$ | 23.36 [-1.14, 47.85] | |
| Fanzania | | |
| Magali KT et al | 42.90 [35.21, 50.59] | 3.98 |
| Heterogeneity: $\tau^2 = 0.00$, $I^2 = .\%$, $H^2 = .$ | 42.90 [35.21, 50.59] | |
| Test of $\theta_i = \theta_j$: Q(0) = 0.00, p = . | • | |
| Jganda | | |
| Holt L et al | 11.00 [8.38, 13.62] | 4.42 |
| Kayondo M et al | 22.10 [12.83, 31.37] | 3.79 |
| Heterogeneity: $\tau^2 = 49.53$, $I^2 = 80.40\%$, $H^2 = 5.10$ Test of $\theta_i = \theta_i$: Q(1) = 5.10, p = 0.02 | 15.62 [4.90, 26.35] | |
| Tambia | | |
| Lamoia Holme A et al | - 27 10 [21 40 22 41 | 4.20 |
| makanda MM at al | 27.10[21.09, 32.51] | 4.22 |
| HIGKAINUU IVINI CI AI Heterongeneity: $r^2 = 25.14$ $\Gamma^2 = 40.30\%$ $\Pi^2 = 1.07$ | 22 96 [14 61 22 103 | 3.25 |
| Test of $\theta_i = \theta_i$: Q(1) = 1.97, p = 0.16 | - 23.80 [14.01, 33.10] | |
| Rwanda | | |
| G. Egziabher T et al | 13.90 [9.79, 18.011 | 4.32 |
| Heterogeneity: $\tau^2 = 0.00$, $I^2 = .\%$, $H^2 = .$ Test of $\theta_i = \theta_j$: Q(0) = -0.00, p = . | 13.90 [9.79, 18.01] | |
| Overall | 24.92 [20.34, 29.50] | |
| Heterogeneity: $\tau^2 = 122.00$, $I^2 = 97.16\%$, $H^2 = 35.26$ | | |
| $\text{fest of } \theta_i = \theta_j; Q(23) = 810.99, p = 0.00$ | | |
| Test of group differences: Q _b (12) = 382.75, p = 0.00 | | |
| 0 20 | 40 60 | |

Random-effects DerSimonian-Laird model

Fig 3. Sub-group analysis of the prevalence of obstetric fistula repair failure in Sub-Saharan Africa based on random effect model, divided by country.

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| Variables | Characteristics | Number of studies | Pooled prevalence (95%CI) | I ² (p-value) |
|--------------------|-----------------|-------------------|---------------------------|--------------------------|
| Publication year | ≤ 2016 | 13 | 21.91 (17.48–26.33) | 94.16% (< 0.001) |
| | > 2016 | 11 | 28.16 (19.21-37.12) | 98.24% (< 0.001) |
| Sample size (mean) | ≤ 411 | 15 | 28.82 (20.77-36.87) | 96.76% (< 0.001) |
| | > 411 | 9 | 18.83 (14.10-23.55) | 96.45% (< 0.001) |

| Table 2. The pooled pr | evalence of obstetric fistula repa | ir failure, 95% CI, and heterog | geneity estimate with a p-value | for the subgroup analysis. |
|------------------------|------------------------------------|---------------------------------|---------------------------------|----------------------------|
|------------------------|------------------------------------|---------------------------------|---------------------------------|----------------------------|

significant publication bias (p-value of 0.021). Moreover, the funnel plot was visually examined for symmetry, and the funnel plot's form also reveals that the asymmetric distribution of the effect estimates is a sign of publication bias (Fig 4). Therefore, the trim and fill analysis was carried out to take publication bias into account.

Trim-and-fill analysis for the prevalence of obstetric fistula repair failure

A trim-and-fill analysis was used to estimate the number of potential missing studies. No studies were imputed during the analysis, and no missing studies were identified. After publication bias was adjusted, the estimated pooled prevalence of obstetric fistula repair failure among women who underwent surgical repair in SSA countries appeared to be 24.92 (95% CI: 20.340–29.505%). This result is similar to the unadjusted pooled prevalence rate of obstetric fistula repair failure, with comparable degrees of heterogeneity across the studies in the random-effects model analysis ($I^2 = 97.16\%$, $p \le 0.001$) (Fig.5).





https://doi.org/10.1371/journal.pone.0295000.g004

Trim and fill analysis for pooled prevalence of obstetric fistula repair failure

Meta-analysis

| Method | Method Pooled | | 95% CI | | Asymptomatic | | |
|--------|---------------|--------|--------|----------|--------------|----|--|
| | Est | Lower | Upper | Z- value | p-value | | |
| Fixed | 18.557 | 17.820 | 19.294 | 49.323 | 0.000 | 24 | |
| random | 24.922 | 20.340 | 29.505 | 10.660 | 0.000 | | |

Test for heterogeneity: Q= 810.986 on 23 degrees of freedom (p= 0.000)

Moment-based estimate of between studies variance = 122.001

Trimming estimator: Linear

Meta-analysis type: Random-effects model

| Iteration | | estimate | Tn | # to trim | diff | |
|-----------|---|----------|-----|-----------|------|--|
| 1 | | 24.922 | 140 | 0 | 300 | |
| 2 | 1 | 24.922 | 140 | 0 | 0 | |

Note: no trimming performed; data unchanged

Filled

Meta-analysis

| Method | Pooled | 95% CI | | Asyn | No of studies | |
|--------|--------|--------|--------|----------|---------------|----|
| | Est | Lower | Upper | Z- value | p-value | |
| Fixed | 18.557 | 17.820 | 19.294 | 49.323 | 0.000 | 24 |
| random | 24.922 | 20.340 | 29.505 | 10.660 | 0.000 | |

Test for heterogeneity: Q= 810.986 on 23 degrees of freedom (p= 0.000)

Moment-based estimate of between studies variance = 122.001

Fig 5. Trim-and-fill analysis for the prevalence of obstetric fistula repair failure among women who underwent surgical repair in Sub-Saharan African countries.

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Meta-regression

A univariate meta-regression was carried out utilizing sample size, publication year, and response rate as factors in order to pinpoint the potential source of heterogeneity. Only sample size, as was found using random-effects meta-regression analysis, was a significant source of heterogeneity. The chance of obstetric fistula repair failure decreased by a factor of 0.014 as the sample sizes increased ($\beta = -0.014$ (95% CI: -0.025-0.004), p-value = 0.007). 24.88% of the variation in overall estimates was explained by univariate meta-regression by sample size (Table 3).

| Table 3. | Univariate meta-regression analysis to determine factors related to the heterogeneity of the prevalence of obstetric fistula repair failures in Sub-Saharan |
|-----------|---|
| African c | ountries. |

| Variables | Coefficient | 95% confidence interval | I ² (%) | R ² (%) | p-value |
|------------------|-------------|-------------------------|--------------------|---------------------------|---------|
| Publication year | 0.311 | - 0.788-1.399 | 97.26 | 0.00 | 0.575 |
| Response rate | 0.158 | - 0.303-0.620 | 97.11 | 0.00 | 0.500 |
| Sample size | - 0.014 | - 0.025-0.004 | 96.57 | 24.88 | 0.007 |

https://doi.org/10.1371/journal.pone.0295000.t003



Fig 6. Sensitivity analysis of prevalence for each study being removed at a time: Prevalence and 95% confidence interval of obstetric fistula repair failure in Sub-Saharan Africa.

Sensitivity analysis

A sensitivity analysis using the random-effects model was conducted to examine the impact of individual studies on the pooled prevalence of obstetric fistula repair failure among women who underwent repair in SSA. The results showed that that no single study had a significant influence on the overall pooled prevalence of obstetric fistula repair failure among women who underwent repair (Fig 6).

The association between urethral damage and obstetric fistula repair failure

Six studies (three from Ethiopia [26, 35, 49], one from Guinea [44], one from the Democratic Republic of the Congo [36], and one from Cameron [50]) were utilised to identify the association between urethral involvement and obstetric fistula failure repair in women who underwent repair in SSA. As a result, the pooled effects of five studies using random effect meta-analysis showed that women with total urethral damage were 3.5 times more likely than those with an intact urethra to undergo obstetric fistula repair failure (OR = 3.50, 95% CI: 2.09, 4.91) (Fig 7).

The association between fistula size and obstetric fistula repair failure

The findings of six studies, two from the Democratic Republic of the Congo [37, 51], two from Ethiopia [32, 35], and two from Uganda [24, 41], respectively, showed that women with large fistula sizes (> 3 cm) were more likely to experience obstetric fistula repair failure than women with small fistula size (less than or equal to 3 cm). Our pooled data revealed that the likelihood of obstetric fistula repair failure was 3 times higher in women with fistulas larger than 3 cm than in those with fistulas smaller than or equal to 3 cm (OR = 3.09, 95% CI: 2.00,4.10), with no heterogeneity among studies (Fig 8).



Fig 7. The pooled effects of complete urethral damage on obstetric fistula repair failure among women who underwent repair in Sub-Saharan Africa.

The association between duration of labor and obstetric fistula repair failure

The relationship between the duration of labor and the failure of obstetric fistula repair was examined in four studies. Two studies from Ethiopia [32, 35] and one from Rwanda [46] found significant associations, but a study from Uganda [24] found no significant relationship between the two variables. According to the pooled analysis, women whose labors lasted less than or equal to 48 hours were 55% less likely to experience obstetric fistula repair failure than those whose labors were longer (> 48 hours) (OR = 2.04, 95% CI: 1.27, 2.81) (Fig 9).

The relationship between previous fistula repair and obstetric fistula repair failure

In order to determine the relationship between prior fistula repair and obstetric fistula repair failure, five studies from Uganda [24, 41], Kenya [27], Rwanda [46], and Democratic republic of the Congo [25] were identified which found a significant association between prior fistula repair and obstetric fistula repair failure. However, a single study conducted in Nigeria revealed no association between prior fistula repair and obstetric fistula repair failure [39]. The results of the pooled data showed that women who had previously had an obstetric fistula repair failure increased likelihood of having an obstetric fistula repair failure (OR = 2.70, 95% CI: 1.94 to 3.45, p-value < 0.05) (Fig 10).



Fig 8. Forest plot for the association between fistula size and obstetric fistula repair failure among women who underwent repair in Sub-Saharan Africa.

The association between postoperative complication and obstetric fistula repair failure

Three studies were identified: one from Uganda, one from Kenya, and one from Ethiopia. Of the three, the Ugandan [24] and Kenyan [27] studies found no significant relationship between post-operative complications and obstetric fistula repair failure. The Ethiopian [32] study found a significant association between postoperative complications and obstetric fistula repair failure. Moreover, the pooled result showed that there is no significant association between postoperative complications and obstetric fistula repair failure. (Fig 11).

Discussion

In accordance with WHO's Sustainable Development Goal 3.7, "by 2030, ensure universal access to sexual and reproductive health-care services, including family planning, information, and education, and the integration of reproductive health into national strategies and programs [52]." Fistula eradication by 2030 is a component of this objective [21]. In order to achieve this objective, women needed to have access to high-quality surgery [53]. Therefore, this systematic review and meta-analysis aimed to assess the pooled prevalence of obstetric fistula repair failure and its associated factors among women who underwent repair in SSA countries.



Fig 9. Forest plot for the association between duration of labor and obstetric fistula repair failure among women who underwent repair in Sub-Saharan Africa.

The results of this meta-analysis indicate that the pooled prevalence of obstetric fistula repair failure in SSA countries was 24.92% [95% CI: 20.34–29.50%]. This is comparable to a multicounty prospective study that revealed a nearly 20% failure rate for obstetric fistula repair [54]. These results were higher than those of a study conducted in Pakistan, which showed that 12.8% of women who underwent obstetric fistula repair experienced failure [55], and a systematic review study that revealed the failure rate after closure was 13% [56]. Additionally, it was higher than the WHO target of less than 15%, which is considered to be a sign of high-quality medical care [1]. There could be a difference in how the failure rate is defined, which could explain the discrepancy between this study's possible differences from previous studies and the WHO target. Other possible reasons for the differences include variations in the type of repair technique, surgical expertise, and the accessibility of high-quality fistula repair services. Failure was measured, as in most studies, two weeks post-surgery [51].

Furthermore, the prevalence of obstetric fistula repair failure is strongly influenced by the patient's age, nutritional state, and body mass index (BMI). In contrast to older patients who may have comorbidities and lower tissue quality, which can raise the chance of repair failure, younger people recover tissue more quickly and are generally in better health. Patients with a history of multiple deliveries may have severe tissue damage and scarring, which makes the repair more difficult and raises the chance of failure. Low BMI can also impact immune system function and tissue healing, increasing the risk of repair failure. Inadequate postoperative care can result in wound breakdown, infection, and other issues that can lead to repair failure. Prolonged obstructed labor during previous deliveries can also cause tissue ischemia and necrosis,



Fig 10. Forest plot for the relationship between previous fistula repair and obstetric fistula repair failure among who women underwent repair in Sub-Saharan Africa.

which further complicates the healing process. Therefore, access to high-quality emergency care, a strong political commitment to effective fistula care, infection prevention, postoperative monitoring, sufficient preoperative nutrition, and BMI optimization are crucial for successful obstetric fistula repair.

The results of the sub-group analysis showed that there were variations in the pooled prevalence of obstetric fistula repair failure among the SSA countries, with Angola having the highest prevalence of 58% [28], followed by a study in Burkina Faso [48], and Rwanda having the lowest prevalence of 13.9% [46]. There could be several reasons for the variations in the rate of obstetric fistula repair failure among the SSA countries, including variations in the accessibility of facilities equipped with the necessary tools and repair techniques, surgical experience and training of the surgeons, fistula characteristics, nutritional status, post-operative nursing care, multiparty, and the presence of comorbid infections. Therefore, implementing interventions suitable for each country is essential to rise the success rate of obstetric fistula repair, particularly in assessing women's nutritional status, any presence of co-morbidities, and enhancing pre- and post-operative care to reduce complications following obstetric fistula repair.

Women with total urethral damage were more likely than those with an intact urethra to experience obstetric fistula repair failure. This finding is in line with a systematic review study [57] and a multicounty prospective study [54]. The finding might be related to the fact that most women who give birth vaginally, have prolonged labors, and have had multiple deliveries are more likely to have a fistula that extends to the urethra. Urethral involvement can also



Fig 11. Forest plot for the association between postoperative complications and obstetric fistula repair failure among women who underwent repair in Sub-Saharan Africa.

affect bladder size and sphincter mechanisms, which can make surgical repair more difficult. Persistent irritation of the wound site from urine incontinence may also raise the risk of repair failure [54]. As a result, for women who benefit from a Caesarean section (CS), the extension of the fistula that involves the urethra is limited compared to those who still deliver vaginally, which might also reduce the problem encountered in prolonged labor.

The likelihood of obstetric fistula repair failure was higher among women with big fistula sizes (> 3 cm) than among those with smaller or equal to 3 cm fistula sizes. This finding was in line with a study on factors determining recurrence after vesicovaginal fistula repair that was undertaken in Pakistan [55] and a systematic review study [58]. However, a multicounty prospective study indicated that there is no relationship between the size of the fistula and obstetric fistula repair failure [54]. The findings might be explained by the larger fistula's potential for difficulty in fully mobilizing to permit a tension-free repair, which could make the procedure more challenging and complex. It may also be related to the idea that the greater the wound edge, the higher the risk of infection and dehiscence [49, 59]. Due to the increased tissue damage and complexity of the treatments involved, larger fistulas have a worse success rate for fistula repair. For repairs to be successful, early discovery and prompt intervention are essential. Addressing this relationship calls for a multidisciplinary strategy that involves specialist surgical knowledge, thorough pre- and post-operative care, and initiatives to provide timely access to obstetric services. By addressing the specific challenges associated with large fistulas, it is possible to improve surgical outcomes and ultimately reduce the burden of obstetric fistulas for affected women.

Women with a history of prior obstetric fistula repair were at a higher risk of experiencing repair failure. These results are consistent with multicounty prospective research [54] and a systematic review study [58]. The findings might be explained by the fact that repeated attempts to close an obstetric fistula result in more vaginal scarring, further harm to healthy tissues, and a foreign body reaction to the stitches, all of which increase the risk of repair failure and decrease the likelihood of regaining physiological function. As a result, women who have experienced several repairs can be more susceptible to health issues such as infection, pain, erectile dysfunction, and secondary infertility [58]. The study emphasizes the importance of selecting operable patients, hiring skilled surgeons, optimizing fistula repair outcomes, and developing national policies. It recommends extended hospital stays for women, scheduling follow-up sessions, and promoting post-repair care [59].

Moreover, women with prolonged labor (> 48 hours) were associated with a higher risk of experiencing obstetric fistula repair failure. The finding could be explained by the fact that ischemia necrosis of the soft tissues around the bladder, rectum, and/or vagina results in tissue damage and associated scarring, which gets worse with the length of labor and ultimately causes repair failure [33, 59]. Furthermore, inadequate nutrition, restricted access to health-care, and postponing seeking medical attention are among the additional risk factors for fistula repair failure that are frequently linked to prolonged labor. These elements may raise the risk of a fistula recurrence and lead to poor surgical results. Obstetric fistulas can have negative psychological and emotional effects in addition to physical ones, which may affect the success of repair surgeries. Improving access to prompt and expert obstetric care, addressing underlying social determinants of health, and offering comprehensive support to women undergoing fistula repair surgeries are all necessary components of a comprehensive strategy to address the link between prolonged labor and obstetric fistula repair failure.

Despite the non-significant association between post-operative complications and obstetric fistula repair failure in this meta-analysis, women who experienced postoperative complications or infections were more likely than their counterparts to encounter fistula repair failure. The quality of services, especially the quality of surgery, is likely to be related to complications such as urinary retention, wound infection, and bleeding. Furthermore, healing might be impaired by an infection at the repair site [60]. Therefore, implementing infection prevention strategies (IPS) or providing postoperative care together with regular patient monitoring at each visit may help lower the likelihood of obstetric fistula repair failure.

Implication for clinical practice

The study highlights the pressing need to enhance fistula repair programs and address the high prevalence of obstetric fistula in SSA. In order to lower the fistula repair failure rate and enhance the quality of life for women, country-specific interventions are required. The study points out the potential and gaps in fistula repair services and emphasizes how crucial it is to fund women's empowerment, health, and rights in order to achieve the Sustainable Development Goals (SDG) [52].

The development of surgical techniques, instruments, and infrastructure is essential for surgical success. Specifically, the study emphasizes the significance of a multi-disciplinary to patient management by evaluating the effectiveness of preoperative and postoperative treatment and assessing women's nutritional status, maternal age, co-morbidity, and obstructed labor as critical risk factors associated with repair failure in women with obstetric fistulas. For obstetric fistula repair to be successful, these issues must be addressed through preoperative optimization, careful surgical technique, and comprehensive postoperative care. It also suggests incorporating these findings into clinical practices to improve women's quality of life and contribute to the global literature.

Limitation of the study

Only English-language articles were taken into consideration for this systematic review and metaanalysis. In addition, the majority of the studies included in the analysis had small sample sizes, which could have an impact on the pooled results. Only 13 SSA countries were represented; as a result, conducting a meta-analysis with a small number of articles reduces statistical power, allows for large standard errors. The causes indicated above led to the discovery of publication bias in this meta-analysis. Only 13 SSA countries were represented in the included research; all of the SSA countries may be underrepresented, due to their poor quality, inadequate description of the outcome variable, and inaccessibility of the full-text article despite repeated requests to the corresponding authors of the original research. Furthermore, as most research reports repair outcomes using different definitions, comparisons of obstetric fistula repair failures become challenging. Others describe obstetric fistula failure as a combination of failed fistula closure and continence after repair surgery, while others merely show the overall rates of fistula closure.

Conclusion and recommendations

Women who received surgical treatment for obstetric fistulas in SSA countries experienced more repair failures than what the WHO recommends. Obstetric fistula repair failure was affected by status of urethral damage, fistula size, duration of labor, types of fistula, and history of previous repairs.

Therefore, policymakers and health care organizations that focus on women's health should prioritize reducing obstetric fistula repair failure and preventing all risk factors (poor nutrition, multiparty, obstructed labor, and maternal age) that lead to conditions like large fistulas, urethral damage, and repeat repair. As a result, country-specific interventions were recommended to improve the application and effectiveness of fistula care initiatives as well as the standard of services at repair facilities.

Expanding sufficient facilities that offers pre- and post-operative care and outfitted with trained surgeons is also essential. These facilities should also handle patient transportation concerns, assess women's nutritional status, and co-morbidity, and show a strong political commitment to efficient fistula care.

Moreover, the current study's findings will contribute to the body of global literature and provide guidance for future research. Specifically, we advise that future investigations combine all studies that have been published in both English and non-English languages and include all qualitative studies in order to more precisely identify the extent of obstetric fistula repair failure in SSA.

Supporting information

S1 File. PRISMA (Preferred Reporting Items for Systematic review and Meta-Analysis) 2020 checklist: An updated guideline for reporting systematic reviews. Recommended items addressed in our systematic review and meta-analysis. (DOCX)

S2 File. Dataset on obstetric fistula repair failure among women who underwent repair in Sub-Saharan Africa.

(XLSX)

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