

### 1039 Appendix S3. Testing model assumptions

1040 In the analysis presented in the main text, we used a Beta(1, 2) prior for the IAR in each  
1041 country or territory model. The intention of this prior was to lightly constrain our ZIKV IAR  
1042 estimates and prevent the model from converging towards extreme estimates without precluding  
1043 the possibility of values anywhere between 0 and 1. This prior distribution has a median value  
1044 of 0.292 (95% range: 0.013–0.842). To examine the sensitivity our IAR estimates to this prior  
1045 assumption, we also ran a model version for each territory with a uniform prior on the IAR. With  
1046 the uniform prior for the subnational IARs, the posterior IAR estimates at both the subnational  
1047 and national level were higher for all 15 modeled countries and territories except for Costa Rica  
1048 (S3 Table; S2 Fig).

1049 In the analysis presented in the main text, we assumed that the reporting of symptomatic  
1050 ZIKV infections,  $\mathcal{Z}_i$ , as suspected cases,  $S_i$ , followed a binomial distribution,  $S_{T,i} \sim \text{Bin}(\mathcal{Z}_i, \rho_{S_{T,i}})$ .  
1051 The probability of a symptomatic infection being reported as a suspected case,  $\rho_{S_{T,i}}$ , in admin-  
1052 istrative unit  $i$  of a country or territory followed a beta distribution with hyperparameters  $\alpha_{S_T}$   
1053 and  $\beta_{S_T}$ . However, because there is considerable overlap between the symptoms of a ZIKV  
1054 infection and the symptoms of several other arbovirus infections—including dengue and chikun-  
1055 gunya—the number of suspected Zika cases could exceed the number of symptomatic ZIKV  
1056 infections if other arbovirus infections were misdiagnosed as ZIKV during the epidemic. To  
1057 account for this possibility, we also considered a model where  $S_{T,i} \sim \text{Poisson}(\mathcal{Z}_i \rho_{S_{T,i}})$ . The  
1058 reporting probability for suspected cases was allowed to range above one by drawing from a  
1059 gamma distribution,  $\rho_{S_{T,i}} \sim \text{Gamma}(\alpha_{S_T}, 1/\beta_{S_T})$ . The gamma distribution hyperparameters,  
1060  $\alpha_{S_T}$  and  $\beta_{S_T}$ , were assigned truncated standard normal prior distributions. These hyperparam-  
1061 eter priors result in a mean of  $\alpha_{S_T} \beta_{S_T} = 0.64$ , and a variance of  $\alpha_{S_T} \beta_{S_T}^2 = 0.512$  for the prior  
1062 of the gamma distribution.

1063 The version of the model with a Poisson distribution for suspected cases was run for Costa  
1064 Rica, Guatemala, Panama, and Puerto Rico. These four countries and territories represented  
1065 the range of estimated suspected reporting probabilities that were observed for the model with  
1066 a binomial distribution, with Guatemala and Panama having relatively low estimates of  $\rho_{S_T}$ ,

1067 and Costa Rica and Puerto Rico having the second highest and highest estimates of  $\rho_{S_T}$ , re-  
1068 spectively. The values of  $\rho_{S_T}$  did not vary significantly between the models with binomial or  
1069 Poisson distributions for Panama (0.092; 95% CrI: 0.024-0.436 vs. 0.074; 95% CrI: 0.021-0.343)  
1070 or Guatemala (0.027; 95% CrI: 0.001-0.194 vs. 0.040; 95% CrI: 0.0022-0.268) (S3 Fig). In addi-  
1071 tion, the IAR estimates for these two countries differed by <1% (S4 Fig). The median estimate  
1072 of  $\rho_{S_T}$  for Costa Rica was lower with the Poisson distribution (0.14; 95% CrI: 0.029 – 0.768) than  
1073 with the binomial distribution (0.255; 95% CrI: 0.037 – 0.908) (S3 Fig). This decrease in the  
1074 estimated reporting probability was only associated with a small increase in the IAR estimate  
1075 from 0.092 (95% CrI: 0.019 – 0.193) to 0.102 (95% CrI: 0.026 – 0.206) (S4 Fig - S5 Fig). Puerto  
1076 Rico had the highest estimate of  $\rho_{S_T}$  when suspected cases were binomially distributed, with  
1077 a median of 0.933 and a upper 95% credible interval very close to 1 (95% CrI: 0.632 – 0.999).  
1078 When we assumed suspected cases followed a Poisson distribution, the median estimate of  $\rho_{S_T}$   
1079 was 0.299 (95% CrI: 0.099 – 0.958). The marginal posterior distribution for  $\rho_{S_T}$  with a Poisson  
1080 distribution was much broader than with the binomial distribution, and although the 95% cred-  
1081 ible interval was below 1 the upper estimates from the posterior were greater than 1 (S3 Fig).  
1082 The estimated territory-wide IAR in Puerto Rico was higher with a Poisson distribution (0.38;  
1083 95% CrI: 0.325 – 0.437) than with the binomial distribution (0.316; 95% CrI: 0.288 – 0.345)  
1084 (S4 Fig). A majority of Puerto Rico’s municipalities had higher IAR estimates with a Poisson  
1085 distribution, but several estimates were lower than they were with a binomial distribution (S6  
1086 Fig).