

### Appendix S3. Testing model assumptions

In the analysis presented in the main text, we used a  $\text{Beta}(1, 2)$  prior for the IAR in each country or territory model. The intention of this prior was to lightly constrain our ZIKV IAR estimates and prevent the model from converging towards extreme estimates without precluding the possibility of values anywhere between 0 and 1. This prior distribution has a median value of 0.292 (95% range: 0.013–0.842). To examine the sensitivity our IAR estimates to this prior assumption, we also ran a model version for each territory with a uniform prior on the IAR. With the uniform prior for the subnational IARs, the posterior IAR estimates at both the subnational and national level were higher for all 15 modeled countries and territories except for Costa Rica (S3 Table; S2 Fig).

In the analysis presented in the main text, we assumed that the reporting of symptomatic ZIKV infections,  $\mathcal{Z}_i$ , as suspected cases,  $S_i$ , followed a binomial distribution,  $S_{T,i} \sim \text{Bin}(\mathcal{Z}_i, \rho_{S_{T,i}})$ . The probability of a symptomatic infection being reported as a suspected case,  $\rho_{S_{T,i}}$ , in administrative unit  $i$  of a country or territory followed a beta distribution with hyperparameters  $\alpha_{S_T}$  and  $\beta_{S_T}$ . However, because there is considerable overlap between the symptoms of a ZIKV infection and the symptoms of several other arbovirus infections—including dengue and chikungunya—the number of suspected Zika cases could exceed the number of symptomatic ZIKV infections if other arbovirus infections were misdiagnosed as ZIKV during the epidemic. To account for this possibility, we also considered a model where  $S_{T,i} \sim \text{Poisson}(\mathcal{Z}_i \rho_{S_{T,i}})$ . The reporting probability for suspected cases was allowed to range above one by drawing from a gamma distribution,  $\rho_{S_{T,i}} \sim \text{Gamma}(\alpha_{S_T}, 1/\beta_{S_T})$ . The gamma distribution hyperparameters,  $\alpha_{S_T}$  and  $\beta_{S_T}$ , were assigned truncated standard normal prior distributions. These hyperparameter priors result in a mean of  $\alpha_{S_T} \beta_{S_T} = 0.64$ , and a variance of  $\alpha_{S_T} \beta_{S_T}^2 = 0.512$  for the prior of the gamma distribution.

The version of the model with a Poisson distribution for suspected cases was run for Costa Rica, Guatemala, Panama, and Puerto Rico. These four countries and territories represented the range of estimated suspected reporting probabilities that were observed for the model with a binomial distribution, with Guatemala and Panama having relatively low estimates of  $\rho_{S_T}$ ,

1067 and Costa Rica and Puerto Rico having the second highest and highest estimates of  $\rho_{S_T}$ , re-  
 1068 spectively. The values of  $\rho_{S_T}$  did not vary significantly between the models with binomial or  
 1069 Poisson distributions for Panama (0.092; 95% CrI: 0.024-0.436 vs. 0.074; 95% CrI: 0.021-0.343)  
 1070 or Guatemala (0.027; 95% CrI: 0.001-0.194 vs. 0.040; 95% CrI: 0.0022-0.268) (S3 Fig). In addi-  
 1071 tion, the IAR estimates for these two countries differed by <1% (S4 Fig). The median estimate  
 1072 of  $\rho_{S_T}$  for Costa Rica was lower with the Poisson distribution (0.14; 95% CrI: 0.029 – 0.768) than  
 1073 with the binomial distribution (0.255; 95% CrI: 0.037 – 0.908) (S3 Fig). This decrease in the  
 1074 estimated reporting probability was only associated with a small increase in the IAR estimate  
 1075 from 0.092 (95% CrI: 0.019 – 0.193) to 0.102 (95% CrI: 0.026 – 0.206) (S4 Fig - S5 Fig). Puerto  
 1076 Rico had the highest estimate of  $\rho_{S_T}$  when suspected cases were binomially distributed, with  
 1077 a median of 0.933 and a upper 95% credible interval very close to 1 (95% CrI: 0.632 – 0.999).  
 1078 When we assumed suspected cases followed a Poisson distribution, the median estimate of  $\rho_{S_T}$   
 1079 was 0.299 (95% CrI: 0.099 – 0.958). The marginal posterior distribution for  $\rho_{S_T}$  with a Poisson  
 1080 distribution was much broader than with the binomial distribution, and although the 95% cred-  
 1081 ible interval was below 1 the upper estimates from the posterior were greater than 1 (S3 Fig).  
 1082 The estimated territory-wide IAR in Puerto Rico was higher with a Poisson distribution (0.38;  
 1083 95% CrI: 0.325 – 0.437) than with the binomial distribution (0.316; 95% CrI: 0.288 – 0.345)  
 1084 (S4 Fig). A majority of Puerto Rico’s municipalities had higher IAR estimates with a Poisson  
 1085 distribution, but several estimates were lower than they were with a binomial distribution (S6  
 1086 Fig).