

# Supporting Information: The inference of sex-biased human demography from whole-genome data

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## Supporting Information

### Simulation Commands

**Population of constant size** We used the program `ms` to simulate data from a population with a sex-biased demographic history. For a population of constant size with no sex-bias ( $p = 0.5$ ) or one with a strong male bias ( $p = 0.2$ ), we used the following commands to simulate autosomal and X-chromosomal data:

```
ms 40 10000 -t 30.0 -r 30.0 5000      # autosomal, p = 0.5
ms 30 10000 -t 22.5 -r 15.0 5000      # X chromosome, p = 0.5
ms 40 10000 -t 19.2 -r 19.2 5000      # autosomal, p = 0.2
ms 30 10000 -t 12.0 -r 4.0 5000       # X chromosome, p = 0.2
```

Analogous commands were used for populations with other demographic histories.

### Likelihood ratio tests for sex-bias: general form

We define a demographic history as a set of population sizes ( $N_{e1}, N_{e2}, \dots, N_{eT}$ ) which go forward in time (i.e.,  $N_{e1}$  is the ancestral population size) and correspond to a set of  $T - 1$  size changes and  $T$  epoch durations. The size changes  $\vec{\nu} = (\nu_1, \nu_2, \dots, \nu_{T-1})$ , which occur instantaneously or exponentially, are defined as the size at the end of an epoch relative to the ancestral population size. The epoch durations  $\vec{\tau} = (\tau_1, \tau_2, \dots, \tau_T)$  are in units of genetic time scaled by the ancestral population size. We assume the X chromosome has the same demographic model (i.e. number and kind of size changes) as the autosomes.

During the epoch  $t$ ,  $t = 1 \dots T$ ,  $N_t^X = c_t \times N_t^A$  for some constant  $c_t$ , which is a function of  $p_t$ , the female fraction of the effective size during the epoch. To test for sex-bias during epoch  $i$ , the constraints  $\nu_t^X = c_{i+1}/c_1 \times \nu_t^A$  and  $\tau_t^X = 1/c_{i+1} \times \tau_t^A$  are used (see derivation for a bottleneck model in “Likelihood ratio tests for sex-bias: bottleneck model” below). Since we define the population-scaled mutation

rate,  $\theta$ , in terms of the ancestral population size, it is constrained by  $c_1$  of the first epoch:  $N_1^X = c_1 \times N_1^A$  and  $\theta^X = c_1 \times \theta^A$ .

## Likelihood ratio tests for sex-bias: bottleneck model

Immediately following a population bottleneck, heterozygosity declines rapidly it reaches a minimum amount due to the reduced population size and then increases slowly with the influx of new mutations [3]. The bottleneck reduces the mean number of alleles at a locus, mainly by removing singletons and doubletons [2], which shifts the mode of the allele frequency distribution toward more common alleles [1]. Bottlenecks affect the X chromosomes and autosomes differently due to their different effective population sizes: in the generations immediately following the end of a bottleneck when a population has recovered to its pre-bottleneck size, the X chromosome will have lost more genetic diversity than the autosomes, and the ratio of their genetic diversity will be less than 0.75. Some time later, the X chromosome will recover genetic diversity faster than the autosomes, and it is possible for the ratio of their genetic diversity to be greater than 0.75 [4].

A bottleneck demographic model has two more parameters than a single size change model, so there are more likelihood ratio tests for a bottleneck than for the single size change model described in the main manuscript, “Sex-bias tests for a two-epoch model”. Using the  $\partial a \partial i$  conventions, the bottleneck parameters are:

- $\nu_B$ : the ratio of the bottleneck and ancient population sizes
- $\nu_F$ : the ratio of the contemporary and ancient population sizes
- $\tau_B$ : the length of bottleneck (in units of  $2 \times N_{ancestral}$  generations)
- $\tau_F$ : time since bottleneck recovery (in units of  $2 \times N_{ancestral}$  generations)

We first fit the parameters of a bottleneck model to the autosomal data. We then fit the following models for X-chromosomal data with the  $\partial a \partial i$  Poisson model so that  $\theta_X$  is an explicit parameter.

### Model 0

**No sex-bias:  $p = 0.5$  for all epochs.** A bottleneck model has three epochs indexed by  $i = 1, 2, 3$ . The equations relating the autosomal and X-chromosomal

parameters are as follows:

$$\begin{aligned}
N_i^X &= 3/4 * N_i^A \\
\nu_B^X &= N_B^X / N_1^X = (3/4 * N_2^A) / (3/4 * N_1^A) = N_2^A / N_1^A = \nu_B^A \\
\nu_F^X &= \nu_F^A \\
\tau_1^A &= \tau_B / N_1^A \\
\tau_1^X &= \tau_B / N_1^X \\
\tau_B &= \tau_1^A * N_1^A = \tau_1^X * N_1^X \\
\tau_1^X &= (\tau_1^A * N_1^A) / N_1^X = (\tau_1^A * N_1^A) / (3/4 * N_1^A) = 4/3 * \tau_1^A
\end{aligned}$$

Rearranging the above, the X-chromosomal parameters are related to the autosomal parameters by the constant  $c = 3/4$  for each of the size changes, which are indexed by  $j = B$  during the bottleneck and  $j = F$  after the bottleneck:

$$\begin{aligned}
\nu_j^X &= \nu_j^A \\
\tau_j^X &= 4/3 * \tau_j^A \\
\theta_1^X &= 3/4 * \theta_1^A
\end{aligned}$$

Since the X-chromosomal and autosomal fold-size changes are constrained to be the equal, X-chromosomal event times are constrained to be  $4/3$  the corresponding autosomal times. There are no free parameters in this model, so the likelihood is evaluated without parameter optimization.

## Model 1

**Constant sex-bias:**  $p$  is a value other than 0.5 and is the same for all epochs. The X-chromosomal and autosomal fold-size changes are constrained to be equal, and the X-chromosomal event times are constrained to be a constant factor  $c$  times the corresponding autosomal times. This constraint parameter is a function of the proportion of females via the reduction factors  $f_A(p) = N_e^A / N$  and  $f_X(p) = N_e^X / N$ :

$$c = f_X(p) / f_A(p)$$

This gives the following constraints:

$$\begin{aligned}
\nu_j^X &= \nu_j^A \\
\tau_j^X &= 1/c * \tau_j^A \\
\theta_1^X &= c * \theta_1^A
\end{aligned}$$

## Model 2

**Sex-biased bottleneck:  $p$  is the same before and after the bottleneck, and differs during the bottleneck.**  $\nu_F^X$  and  $\nu_F^A$  after the bottleneck are the same because the X chromosome and the autosomes undergo the same size change and the proportion of females for the first and last epochs are the same. There are two constraint parameters,  $c_1$  and  $c_2$ , which are free parameters:

$$\begin{aligned} c_1 &= f_X(p_1)/f_A(p_1) \\ c_2 &= f_X(p_2)/f_A(p_2) \end{aligned}$$

This gives the following constraints:

$$\begin{aligned} \nu_B^X &= c_2/c_1 * \nu_B^A \\ \nu_F^X &= \nu_F^A \\ \tau_B^X &= 1/c_2 * \tau_B^A \\ \tau_F^X &= 1/c_1 * \tau_F^A \\ \theta_1^X &= c_1 * \theta_1^A \end{aligned}$$

## Model 3

**Changing sex-bias:  $p$  differs for each epoch.** There are three constraint parameters,  $c_1$ ,  $c_2$ ,  $c_3$ , which are free parameters:

$$\begin{aligned} c_1 &= f_X(p_1)/f_A(p_1) & (1) \\ c_2 &= f_X(p_2)/f_A(p_2) & (2) \\ c_3 &= f_X(p_3)/f_A(p_3) & (3) \end{aligned}$$

This gives the following constraints:

$$\begin{aligned} \nu_B^X &= c_2/c_1 * \nu_B^A & (4) \\ \nu_F^X &= c_3/c_2 * \nu_F^A & (5) \\ \tau_i^X &= 1/c_2 * \tau_i^A & (6) \\ \tau_i^X &= 1/c_3 * \tau_i^A & (7) \\ \theta_1^X &= c_1 * \theta_1^A & (8) \end{aligned}$$

These models and their corresponding likelihood ratio tests are used to test for sex-bias in a bottleneck model. Analogous models can be defined and used to test for sex-bias in other demographic models.

## References

- [1] G Luikart, F W Allendorf, and W B Sherwin. Distortion of allele frequency distributions provides a test for recent population bottlenecks. *Journal of Heredity*, pages 238–247, 1998.

- [2] Takeo Maruyama and Paul Fuerst. Population bottlenecks and nonequilibrium models in population genetics. II. Number of alleles in a small population that was formed by a recent bottleneck. *Genetics*, pages 675–689, 1985.
- [3] Masatoshi Nei, Takeo Maruyama, and Ranajit Chakraborty. The Bottleneck Effect and Genetic Variability in Populations. *Evolution*, 29(1):1–10, 1975.
- [4] John E Pool and Rasmus Nielsen. Population size changes reshape genomic patterns of diversity. *Evolution*, 61(12):3001–6, dec 2007.